

# Machine Learning II

DS 4420 - Spring 2020

MLE, MAP, &  
Graphical models

Byron C. Wallace



# Probability Spaces

**Definition:** A probability space  $(\Omega, \mathcal{F}, P)$  consists of

- A sample space  $\Omega$  (i.e. the set of *outcomes*)
- A set of events  $\mathcal{F}$  (i.e. the set possible sets)
- A probability measure  $P$  (maps events to probabilities)

## Axioms of Probability

$$P : \mathcal{F} \rightarrow \mathbb{R} \quad P(E) \geq 0 \quad \forall E \in \mathcal{F} \quad P(\Omega) = 1$$

$$P(E_1, E_2) = P(E_1) + P(E_2) \text{ when } E_1 \cap E_2 = \emptyset$$

# Conditional Probabilities

- **Definition:** Joint Probability

$$P(A, B) = P(A \cap B)$$

Outcomes in both  $A$  and  $B$

Events (i.e. sets of outcomes)

- **Definition:** Conditional Probability

$$P(A | B) = \frac{P(A, B)}{P(B)}$$

# Probability Density Functions

- **Problem:** If  $X$  is a *continuous* variable, then  $P(X=x)$  is 0 for any outcome  $x$

$$X \sim \text{Normal}(0, 1)$$

Single Outcome

$$P(X = \pi) = 0$$

Event

$$P(3.1 \leq X \leq 3.2) \neq 0$$

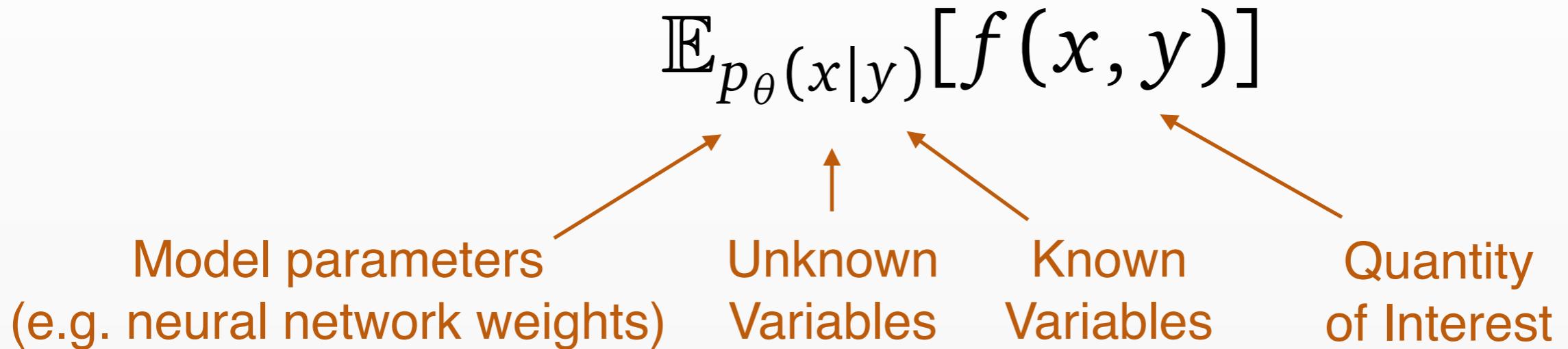
- **Solution:** Define a density function as a derivative

Capital P for probability

$$p_X(x) = \lim_{\delta \rightarrow 0} \frac{P(x - \delta < X < x + \delta)}{2\delta}$$

Small p for density

# Objectives in Learning



Setting	<b>Self-driving Cars</b>	<b>Medical Diagnosis</b>
$p_\theta(y, x)$	Model for pedestrian behavior	Model for diseases / symptoms
$y$	Pedestrian motion	Symptoms / Test results
$x$	Will pedestrian cross road?	Condition of patient
$f(y, x)$	Chance of accident	Treatment outcome

# Maximum Likelihood Estimation

# MLE Framework

Observe some data  $X = x_1, \dots, x_n \quad x_i \in R^d$

We assume this is a random draw (sample)  
from some parameterized distribution  $P_\theta$

# MLE Framework

Observe some data  $X = x_1, \dots, x_n \quad x_i \in R^d$

We assume this is a random draw (sample)  
from some parameterized distribution  $P_\theta$

Goal: find  $\theta$

# MLE Framework

Observe some data  $X = x_1, \dots, x_n \quad x_i \in R^d$

We assume this is a random draw (sample)  
from some parameterized distribution  $P_\theta$

Goal: find  $\theta$

In MLE we pick

$$\theta_{\text{MLE}} = \operatorname{argmax}_\theta P(X|\theta)$$

# MLE Framework

Observe some data  $X = x_1, \dots, x_n \quad x_i \in R^d$

We assume this is a random draw (sample)  
from some parameterized distribution  $P_\theta$

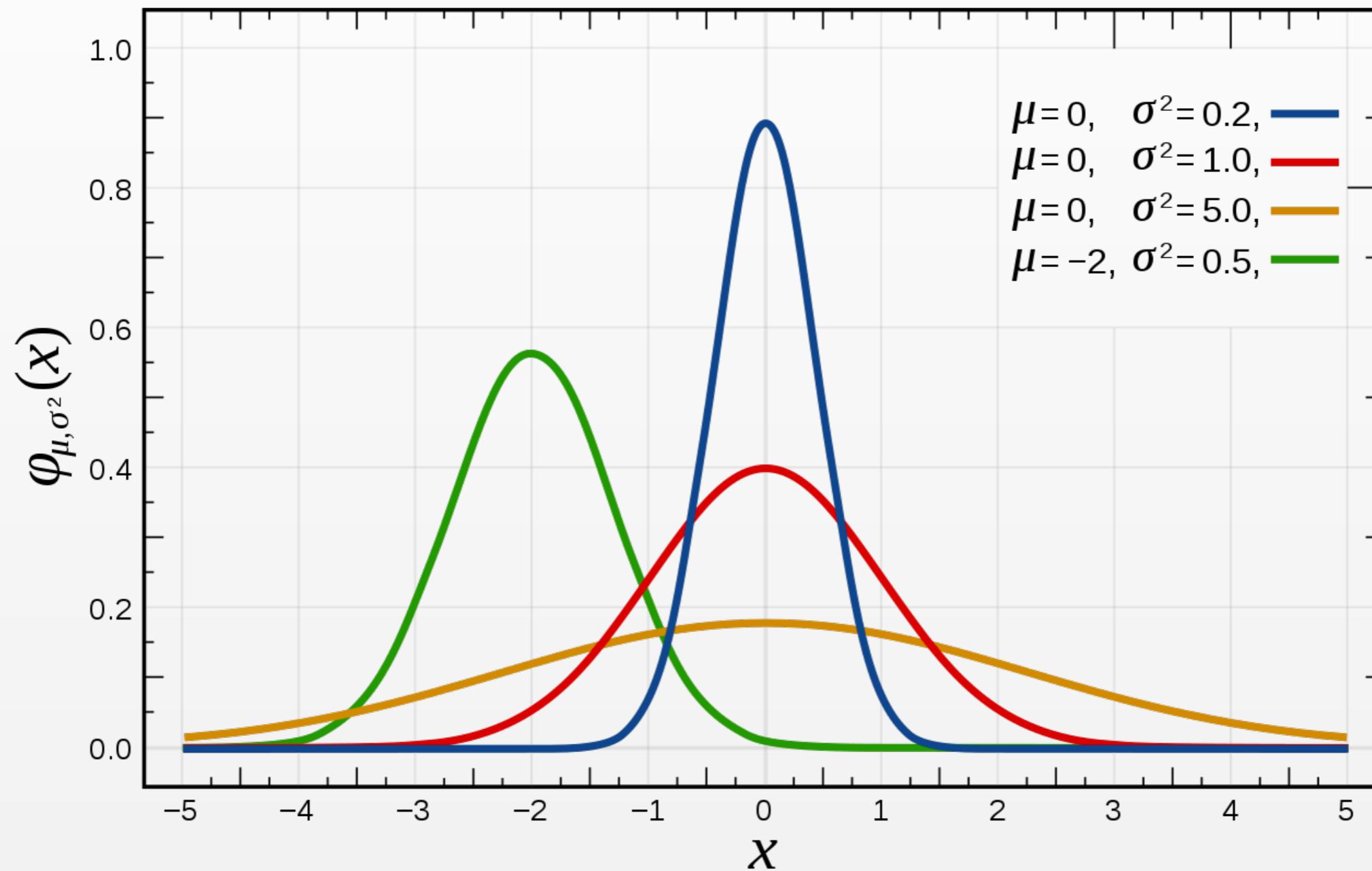
Goal: find  $\theta$

In MLE we pick

$$\theta_{\text{MLE}} = \operatorname{argmax}_\theta P(X|\theta)$$

$$P(X|\theta) = \prod_i P(x_i|\theta)$$

# Normal



$$x \sim N(\mu,\sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}}\exp\{-\frac{1}{2\sigma^2}(x-\mu)^2\}$$

$$\theta=\{\mu,\sigma^2\}$$

$$x \sim N(\mu,\sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}}\text{exp}\{-\frac{1}{2\sigma^2}(x-\mu)^2\}$$

$$\theta=\{\mu,\sigma^2\}$$

$$p(D|\theta)=p(x_1,...,x_N)=\prod_{i=1}^N p(x_i|\theta)$$

$$x \sim N(\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{1}{2\sigma^2}(x - \mu)^2\right\}$$

$$\theta = \{\mu, \sigma^2\}$$

$$p(D|\theta) = p(x_1, \dots, x_N) = \prod_{i=1}^N p(x_i|\theta)$$

*Let's work this out...*

# Discrete/Categorical Distribution

Example: Loaded Dice



$$p_{\theta}(x = k) = \theta_k \quad x \in \{1, 2, 3, 4, 5, 6\}$$
$$\theta = \{\theta_1, \dots, \theta_6\}$$

# Discrete/Categorical Distribution

Example: Loaded Dice



$$p_{\theta}(x = k) = \theta_k \quad x \in \{1, 2, 3, 4, 5, 6\}$$
$$\theta = \{\theta_1, \dots, \theta_6\}$$

Equivalent Notation: Indicator variables

$$p_{\theta}(x) = \prod_{k=1}^K \theta_k^{x_k} \quad x_k := I[x = k]$$

[1, 0, 0, 0, 0, 0] : 1

[0, 1, 0, 0, 0, 0] : 2

...

[0, 0, 0, 0, 0, 1] : 6

# Discrete/Categorical Distribution

Example: Loaded Dice



$$p_{\theta}(x = k) = \theta_k \quad x \in \{1, 2, 3, 4, 5, 6\}$$
$$\theta = \{\theta_1, \dots, \theta_6\}$$

Equivalent Notation: Indicator variables

$$p_{\theta}(x) = \prod_{k=1}^K \theta_k^{x_k} \quad x_k := I[x = k]$$

*Question:* If you perform 1000 rolls and get 200 outcomes  $x=6$ , then how would you estimate  $\theta_6$ ?

$[1, 0, 0, 0, 0, 0] : 1$

$[0, 1, 0, 0, 0, 0] : 2$

...

$[0, 0, 0, 0, 0, 1] : 6$

# Maximum Likelihood Estimation

Likelihood of  $N$  independent events:



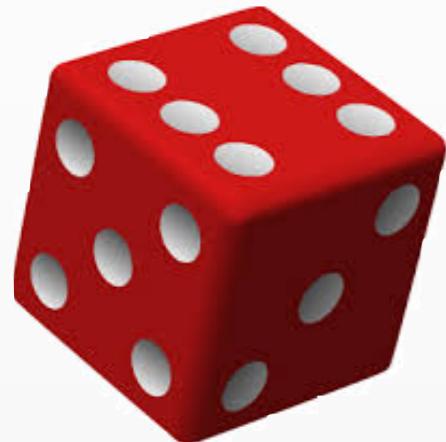
$$p_{\theta}(x_1, \dots, x_N) = \prod_{n=1}^N p_{\theta}(x_n) \quad p_{\theta}(x_n) = \prod_{k=1}^K \theta_k^{x_{n,k}}$$

Maximum likelihood estimation

$$\begin{aligned}\theta^* &= \operatorname{argmax}_{\theta} p_{\theta}(x_1, \dots, x_N) \\ &= \operatorname{argmax}_{\theta} \log p_{\theta}(x_1, \dots, x_N)\end{aligned}$$

# Maximum Likelihood Estimation

Likelihood of  $N$  independent events:



$$p_{\theta}(x_1, \dots, x_N) = \prod_{n=1}^N p_{\theta}(x_n) \quad p_{\theta}(x_n) = \prod_{k=1}^K \theta_k^{x_{n,k}}$$

Maximum likelihood estimation

$$\theta^* = \operatorname{argmax}_{\theta} p_{\theta}(x_1, \dots, x_N)$$

$$= \operatorname{argmax}_{\theta} \log p_{\theta}(x_1, \dots, x_N)$$

*Problem:* Express  $\theta^*$  in terms of  $\{x_1, \dots, x_N\}$

*hint:* solve for  $\nabla_{\theta} \log p_{\theta}(x_1, \dots, x_N) = 0$

# Maximum Likelihood Estimation

Likelihood of  $N$  independent events:



$$p_{\theta}(x_1, \dots, x_N) = \prod_{n=1}^N p_{\theta}(x_n) \quad p_{\theta}(x_n) = \prod_{k=1}^K \theta_k^{x_{n,k}}$$

Maximum likelihood estimation

$$\theta^* = \operatorname{argmax}_{\theta} p_{\theta}(x_1, \dots, x_N)$$

$$= \operatorname{argmax}_{\theta} \log p_{\theta}(x_1, \dots, x_N)$$

$$= \operatorname{argmax}_{\theta} \sum_{k=1}^K N_k \log \theta_k \quad N_k = \sum_{n=1}^N x_{n,k}$$

(known as cross-entropy loss in neural net libraries)

# Likelihood: Bernoulli

(a discrete distribution with outcomes  $x=0$  and  $x=1$ )

$$\begin{aligned}\text{Bern}(x|\mu) &= \mu^x(1-\mu)^{1-x} & \theta_1^{x_1} \theta_2^{x_2} \\ \mathbb{E}[x] &= \mu & x_2 = (1 - x_1) \\ \text{var}[x] &= \mu(1-\mu) & \theta_2 = (1 - \theta_1)\end{aligned}$$

$$\text{mode}[x] = \begin{cases} 1 & \text{if } \mu \geq 0.5, \\ 0 & \text{otherwise} \end{cases}$$

$$x \in \{0, 1\} \quad \mu \in [0, 1]$$

# Likelihood: Bernoulli

(a discrete distribution with outcomes  $x=0$  and  $x=1$ )

$$\text{Bern}(x|\mu) = \mu^x(1-\mu)^{1-x}$$

**Question:** What is the likelihood of  $N$  trials?  
What is ML estimate for  $\mu$ ?

# Likelihood: Bernoulli

(a discrete distribution with outcomes  $x=0$  and  $x=1$ )

$$\text{Bern}(x|\mu) = \mu^x(1-\mu)^{1-x}$$

**Question:** What is the likelihood of  $N$  trials?  
What is ML estimate for  $\mu$ ?

***Let's work this out in the MLE framework***

# Likelihood: Bernoulli

(a discrete distribution with outcomes  $x=0$  and  $x=1$ )

$$\text{Bern}(x|\mu) = \mu^x(1-\mu)^{1-x}$$

**Question:** What is the likelihood of  $N$  trials?

What is ML estimate for  $\mu$ ?

$$p(x_1, \dots, x_N | \mu) = \mu^{N_1} (1-\mu)^{N_0}$$

$$\underset{\mu}{\operatorname{argmax}} p(x | \mu) = \frac{N_1}{N}$$

$$N_1 = \sum_{n=1}^N x_n$$

$$N_0 = \sum_{n=1}^N (1 - x_n)$$

# Problems with MLE?

- Provides a *point estimate*; no notion of uncertainty around parameters
- Does not naturally incorporate prior beliefs (maybe a pro, if you're a frequentist?)
- Other thoughts?

# Maximum A Posteriori (MAP)

*Problem:* Maximum likelihood can overfit when data is limited

*One Solution:* Place a prior over parameters

$$\begin{aligned}\theta^* &= \operatorname{argmax}_{\theta} \log p(\theta | x) \\ &= \operatorname{argmax}_{\theta} \log[p(\theta | x)p(x)] \\ &= \operatorname{argmax}_{\theta} \log[p(x | \theta)p(\theta)] \\ &= \operatorname{argmax}_{\theta} \log p(x | \theta) + \log p(\theta)\end{aligned}$$

ML Objective  
(same as minimizing CE loss)  Regularization

# Intuition: Bayesian Posterior



$$p(\theta | x) = p(x | \theta)p(\theta)/p(x)$$

└ Posterior ┌ Likelihood ┌ Prior

*Example: Biased Coin*



$x$       Observed data (flip outcomes)

$x = 1$ : Heads

$x = 0$ : Tails

$\theta$       Unknown variable (coin bias)

$\theta = 0.5$ : No bias

$\theta = 1.0$ : Always heads

$\theta = 0.0$ : Always tails

# Intuition: Bayesian Posterior



$$p(\theta | x) = p(x | \theta)p(\theta)/p(x)$$

└ Posterior └ Likelihood └ Prior



Uninformative Prior

$p(\theta)$

0 heads, 0 tails

.0    0.2    0.4    0.6    0.8    1.0

$\theta$  (Coin Bias)

# Intuition: Bayesian Posterior



$$p(\theta | x) = p(x | \theta)p(\theta)/p(x)$$

Posterior

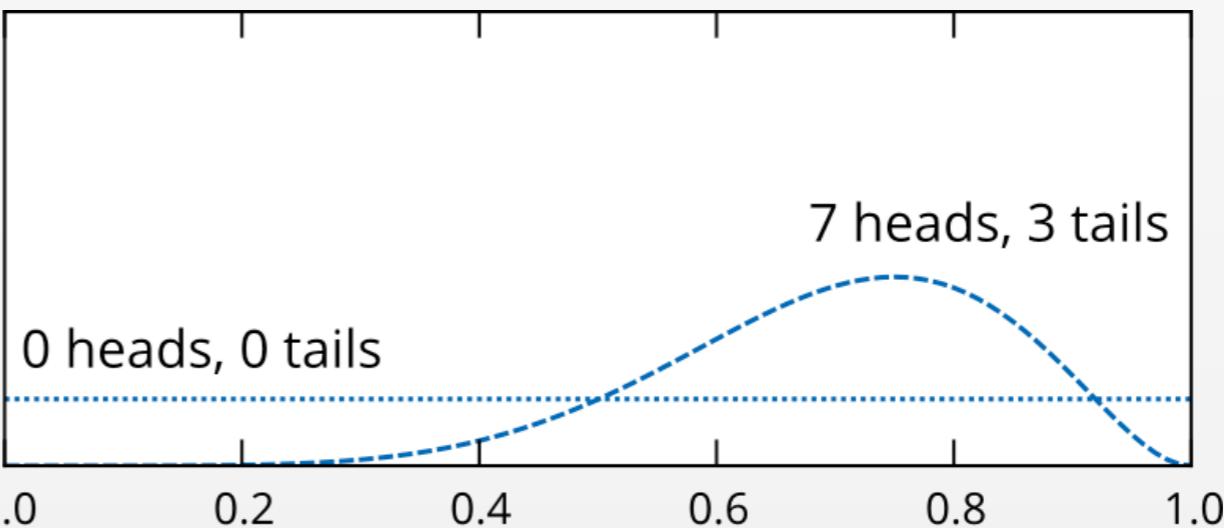
Likelihood

Prior

Posterior after 10 trials



$$p(\theta | x)$$



$\theta$  (Coin Bias)

# Intuition: Bayesian Posterior



$$p(\theta | x) = p(x | \theta)p(\theta)/p(x)$$

Posterior

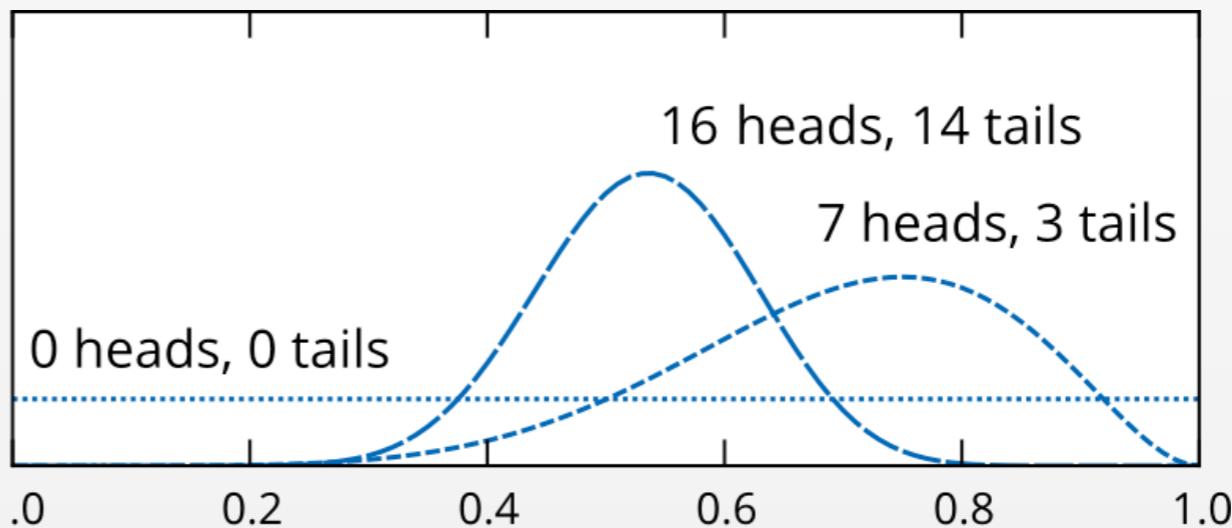
Likelihood

Prior

Posterior after 30 trials



$$p(\theta | x)$$



$\theta$  (Coin Bias)

# Intuition: Bayesian Posterior



$$p(\theta | x) = p(x | \theta)p(\theta)/p(x)$$

Posterior

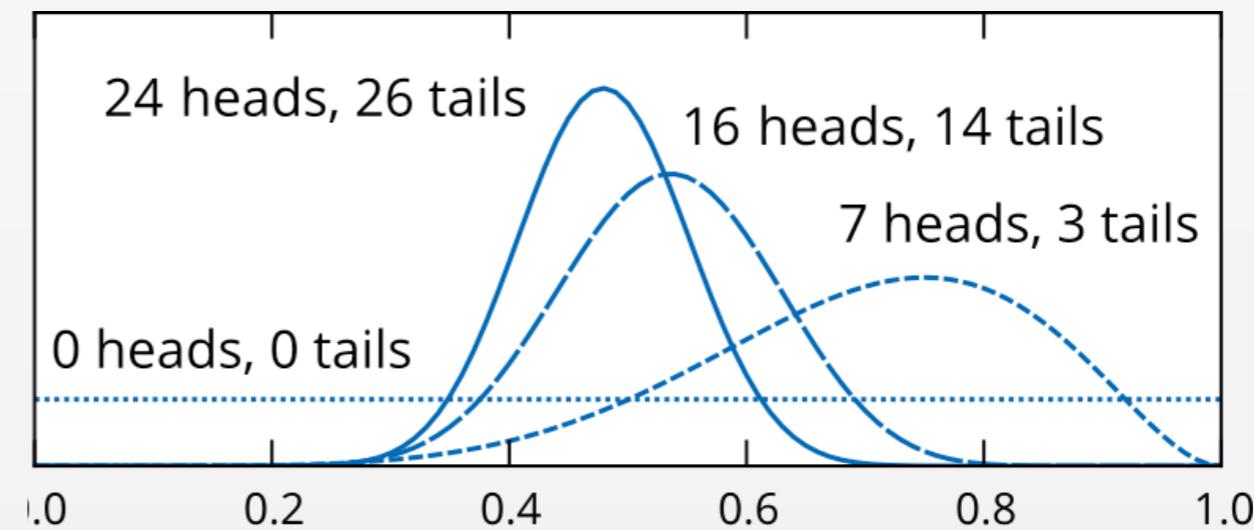
Likelihood

Prior

Posterior after 50 trials



$$p(\theta | x)$$



$\theta$  (Coin Bias)

# Intuition: Bayesian Posterior



$$p(\theta | x) = p(x | \theta)p(\theta)/p(x)$$

Posterior

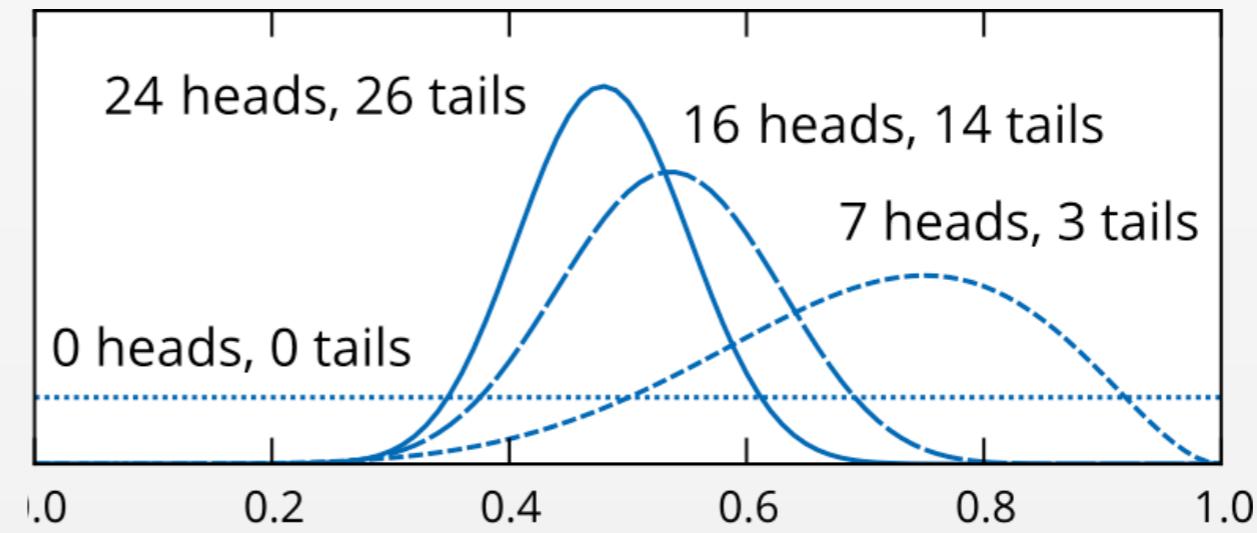
Likelihood

Prior

Posterior after 50 trials



$$p(\theta | x)$$



$\theta$  (Coin Bias)

# Intuition: Bayesian Posterior



$$p(\theta | x) = p(x | \theta)p(\theta)/p(x)$$

Posterior

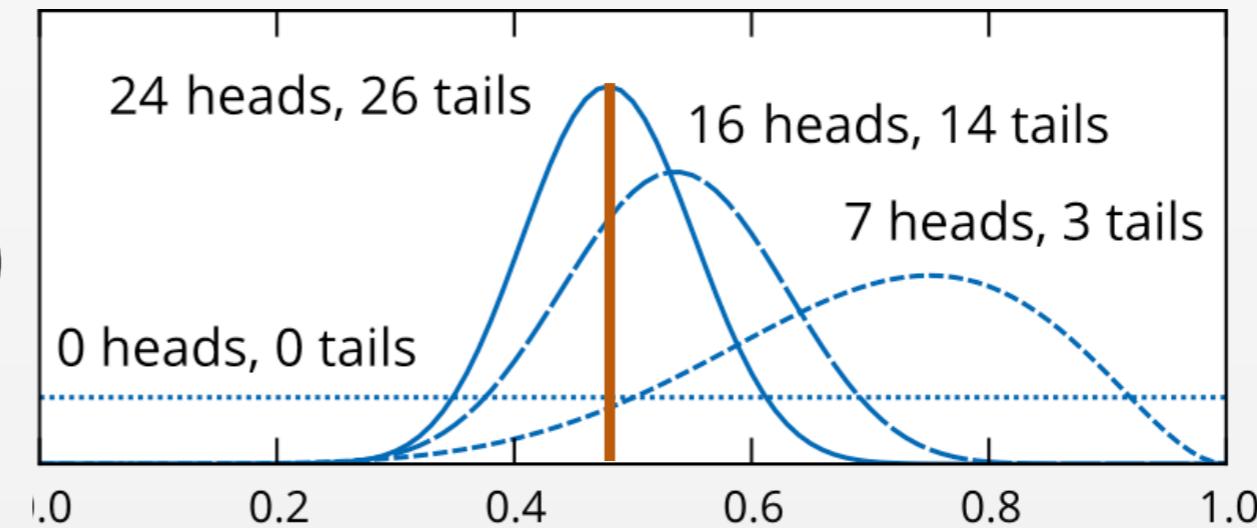
Likelihood

Prior

MAP estimate after 50 trials



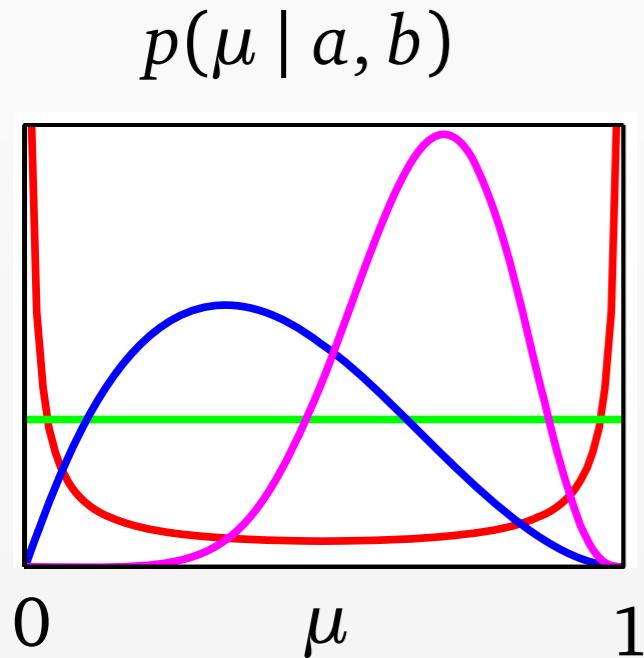
$$p(\theta | x)$$



$\theta$  (Coin Bias)

# Prior: Beta Distribution

(a distribution on values in the range 0 to 1)



$$\begin{aligned}\text{Beta}(\mu|a, b) &= \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)}\mu^{a-1}(1-\mu)^{b-1} \\ \mathbb{E}[\mu] &= \frac{a}{a+b} \\ \text{var}[\mu] &= \frac{ab}{(a+b)^2(a+b+1)} \\ \text{mode}[\mu] &= \frac{a-1}{a+b-2}.\end{aligned}$$

# ML and MAP estimation

Maximum Likelihood Estimation

$$\begin{aligned}\theta^* &= \operatorname{argmax}_{\theta} p_{\theta}(x_1, \dots, x_N) \\ &= \operatorname{argmax}_{\theta} \log p_{\theta}(x_1, \dots, x_N)\end{aligned}$$

Maximum A Posterior Estimation

$$\begin{aligned}\theta^* &= \operatorname{argmax}_{\theta} \log p(\theta | x) \\ &= \operatorname{argmax}_{\theta} \log p(x | \theta) + \log p(\theta)\end{aligned}$$

ML Objective  
(same as minimizing a loss function)

Regularization

The diagram consists of two orange arrows. One arrow points from the text 'ML Objective' to the term  $\log p(x | \theta)$  in the second equation. Another arrow points from the text 'Regularization' to the term  $\log p(\theta)$  in the same equation.

# The Marginal Likelihood

Marginal likelihood  
(a.k.a. the Evidence)

Likelihood      Prior

$$\begin{aligned} p(x | a, b) &= \int p(x | \mu) p(\mu | a, b) d\mu \\ &= \mathbb{E}_{p(\mu|a,b)}[p(x | \mu)] \end{aligned}$$

“Average” likelihood of data (under prior)

$$p(x | a, b) = \frac{B(N_1+a, N_0+b)}{B(a, b)}$$

Can calculate in closed form for conjugate priors

Let's see this in action...

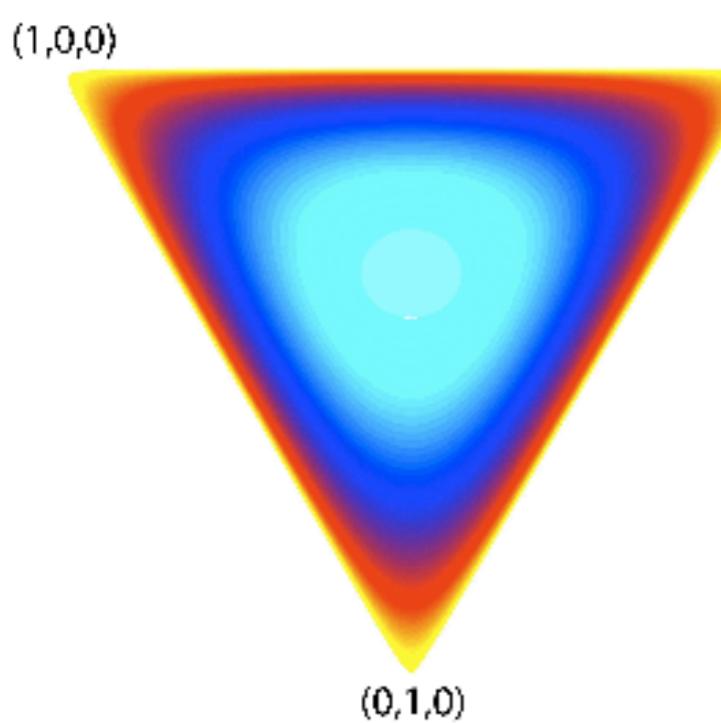
# Example: Dirichlet Distribution

(Conjugate to the Discrete Distribution)

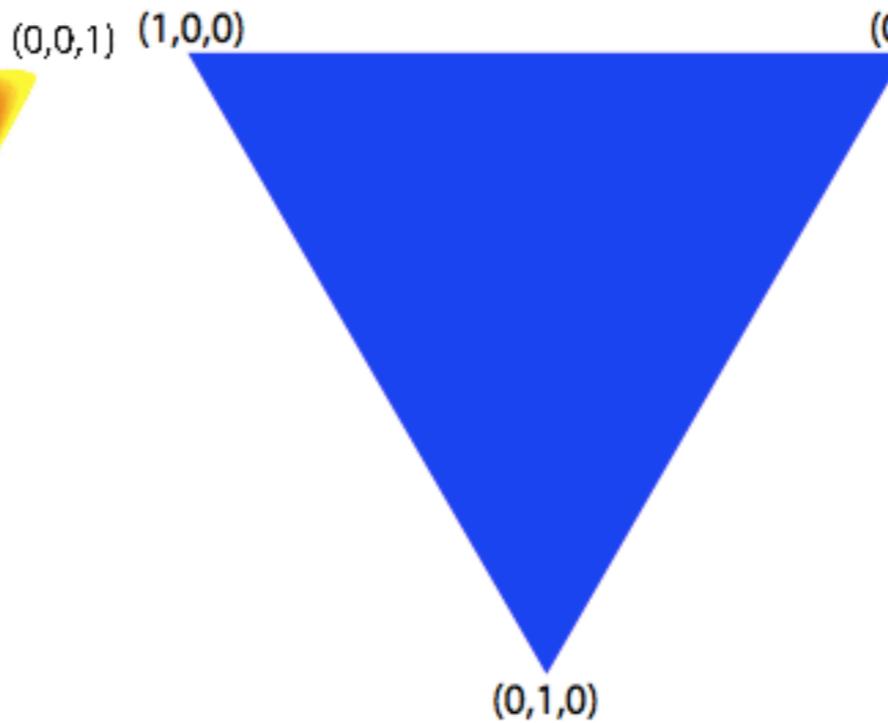
$$p(\theta) = \frac{1}{B(\alpha)} \prod_{k=1}^K \theta_k^{\alpha_k - 1}$$

$$B(\alpha) := \frac{\prod_{k=1}^K \Gamma(\alpha_k)}{\Gamma\left(\sum_{k=1}^K \alpha_k\right)}$$

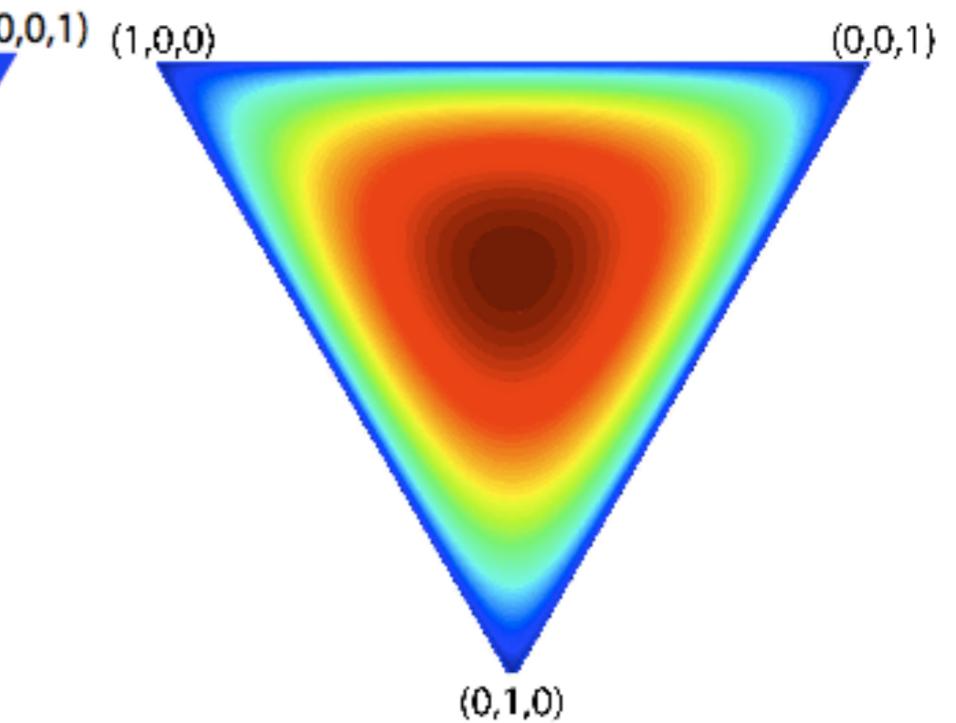
$$\alpha = (0.1, 0.1, 0.1)$$



$$\alpha = (1.0, 1.0, 1.0)$$

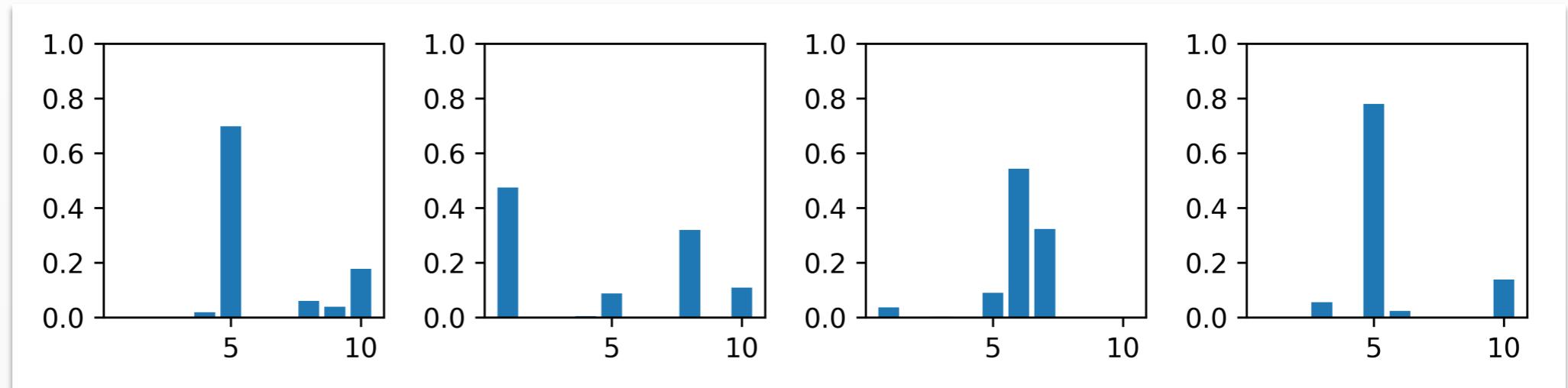


$$\alpha = (10.0, 10.0, 10.0)$$

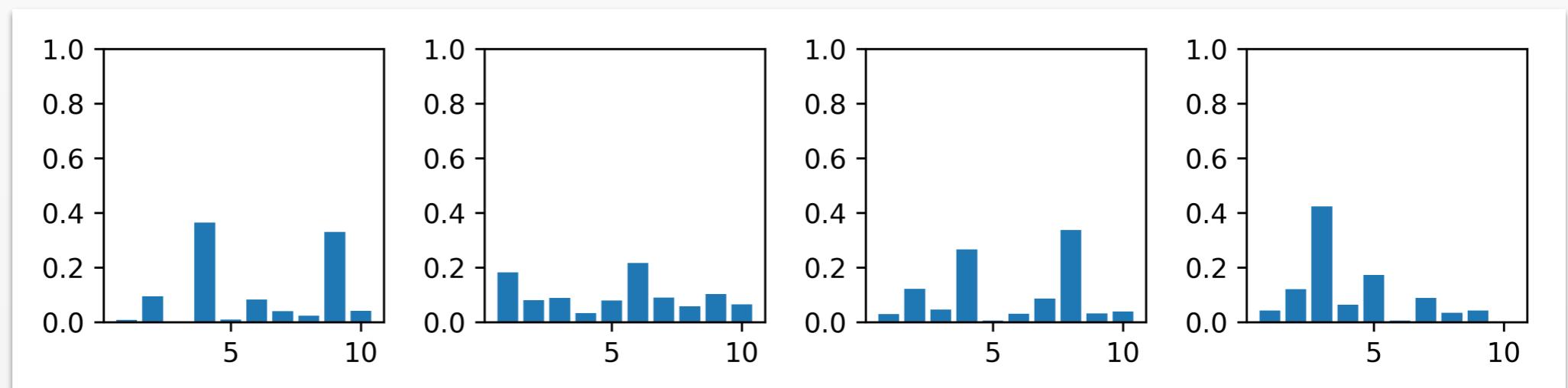


# Example: Dirichlet Distribution

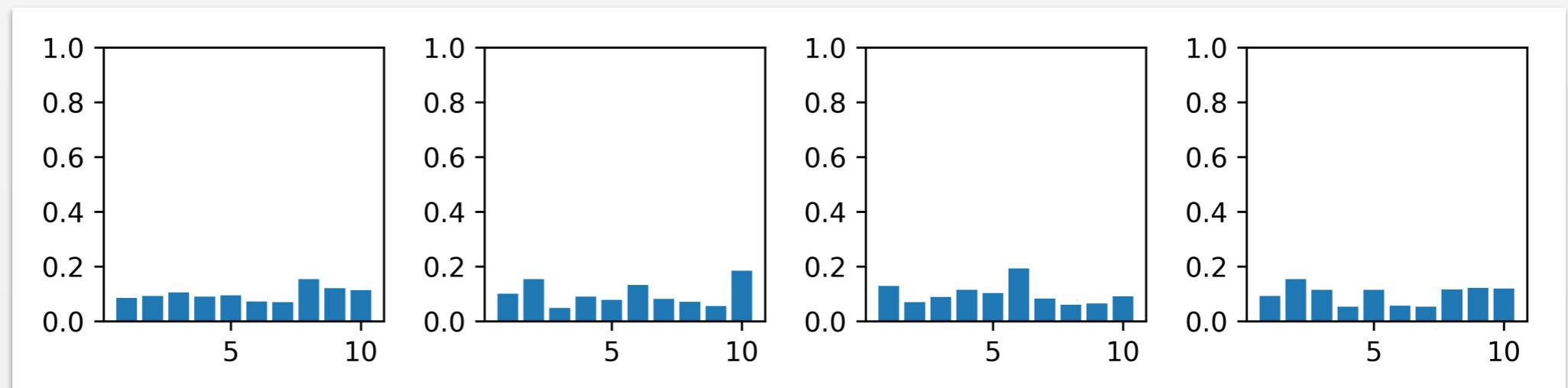
$\alpha_k = 0.1$



$\alpha_k = 1.0$



$\alpha_k = 10.0$



# Conjugate Priors

[https://en.wikipedia.org/wiki/Conjugate\\_prior](https://en.wikipedia.org/wiki/Conjugate_prior)

Likelihood	Model parameters	Conjugate prior distribution	Prior hyperparameters	Posterior hyperparameters	Interpretation of hyperparameters	Posterior predictive <sup>[note 4]</sup>
Normal with known variance $\sigma^2$	$\mu$ (mean)	Normal	$\mu_0, \sigma_0^2$	$\frac{1}{\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}} \left( \frac{\mu_0}{\sigma_0^2} + \frac{\sum_{i=1}^n x_i}{\sigma^2} \right), \left( \frac{1}{\sigma_0^2} + \frac{n}{\sigma^2} \right)^{-1}$	mean was estimated from observations with total precision (sum of all individual precisions) $1/\sigma_0^2$ and with sample mean $\mu_0$	$\mathcal{N}(\tilde{x}   \mu'_0, \sigma_0'^2 + \sigma^2)$ <sup>[5]</sup>
Normal with known precision $\tau$	$\mu$ (mean)	Normal	$\mu_0, \tau_0$	$\frac{\tau_0 \mu_0 + \tau \sum_{i=1}^n x_i}{\tau_0 + n\tau}, \tau_0 + n\tau$	mean was estimated from observations with total precision (sum of all individual precisions) $\tau_0$ and with sample mean $\mu_0$	$\mathcal{N}(\tilde{x}   \mu'_0, \frac{1}{\tau'_0} + \frac{1}{\tau})$ <sup>[5]</sup>
Normal with known mean $\mu$	$\sigma^2$ (variance)	Inverse gamma	$\alpha, \beta$ <sup>[note 5]</sup>	$\alpha + \frac{n}{2}, \beta + \frac{\sum_{i=1}^n (x_i - \mu)^2}{2}$	variance was estimated from $2\alpha$ observations with sample variance $\beta/\alpha$ (i.e. with sum of squared deviations $2\beta$ , where deviations are from known mean $\mu$ )	$t_{2\alpha'}(\tilde{x}   \mu, \sigma^2 = \beta'/\alpha')$ <sup>[5]</sup>
Normal with known mean $\mu$	$\sigma^2$ (variance)	Scaled inverse chi-squared	$\nu, \sigma_0^2$	$\nu + n, \frac{\nu \sigma_0^2 + \sum_{i=1}^n (x_i - \mu)^2}{\nu + n}$	variance was estimated from $\nu$ observations with sample variance $\sigma_0^2$	$t_{\nu'}(\tilde{x}   \mu, \sigma_0'^2)$ <sup>[5]</sup>
Normal with known mean $\mu$	$\tau$ (precision)	Gamma	$\alpha, \beta$ <sup>[note 3]</sup>	$\alpha + \frac{n}{2}, \beta + \frac{\sum_{i=1}^n (x_i - \mu)^2}{2}$	precision was estimated from $2\alpha$ observations with sample variance $\beta/\alpha$ (i.e. with sum of squared deviations $2\beta$ , where deviations are from known mean $\mu$ )	$t_{2\alpha'}(\tilde{x}   \mu, \sigma^2 = \beta'/\alpha')$ <sup>[5]</sup>

# Graphical Models

# Motivation: Spam Filtering

**Features: Words in E-mail**

$x = \begin{bmatrix} 1 \\ 0 \\ 0 \\ \vdots \\ 1 \\ \vdots \\ 0 \end{bmatrix}$	a
	aardvark
	aardwolf
	:
	buy
	:
	zygmurgy

**Labels: Spam or not Spam**

$$y_n \in \{0, 1\}$$

**Input: Labeled Data**

$$\{(x_1, y_1), \dots, (x_N, y_N)\}$$

**Goal: Predict Unlabeled Data**

# Naive Bayes (on board)

# Graphical Model: Naive Bayes

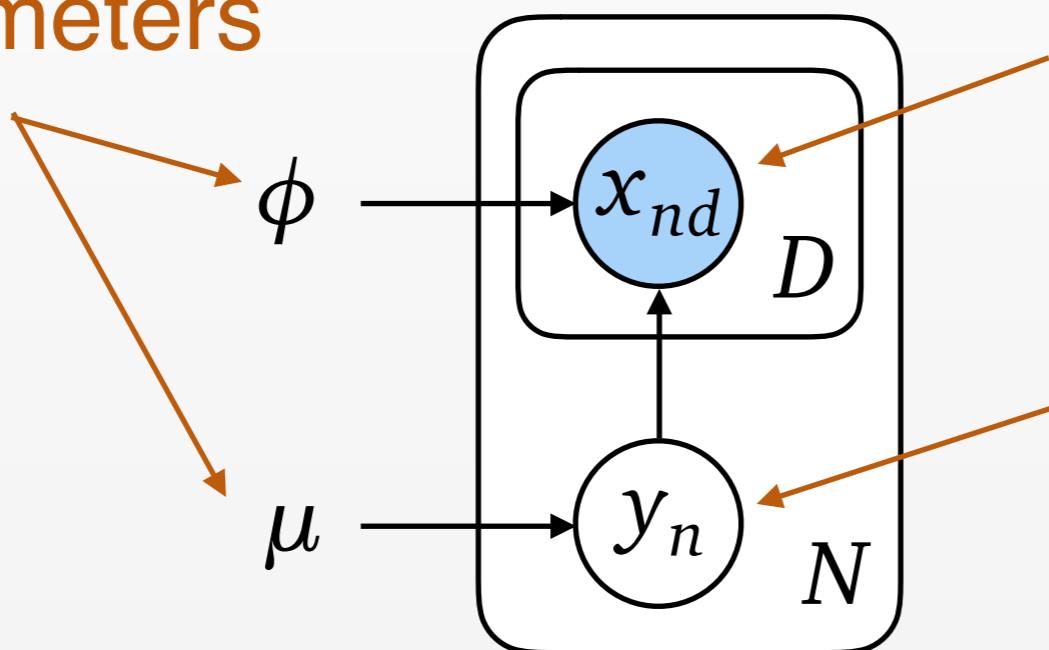
$$y_n \sim \text{Bernoulli}(\mu)$$

$$n = 1, \dots, N$$

$$x_{nd} | y_n = k \sim \text{Bernoulli}(\phi_{kd})$$

$$k = 0, 1 \quad d = 1, \dots, D$$

Parameters



Observed Variables  
(value known)

Unobserved Variables  
(value unknown)

$$p(x, y | \mu, \phi) = \prod_{n=1}^N p(y_n | \mu) \prod_{d=1}^D p(x_{nd} | y_n, \phi)$$

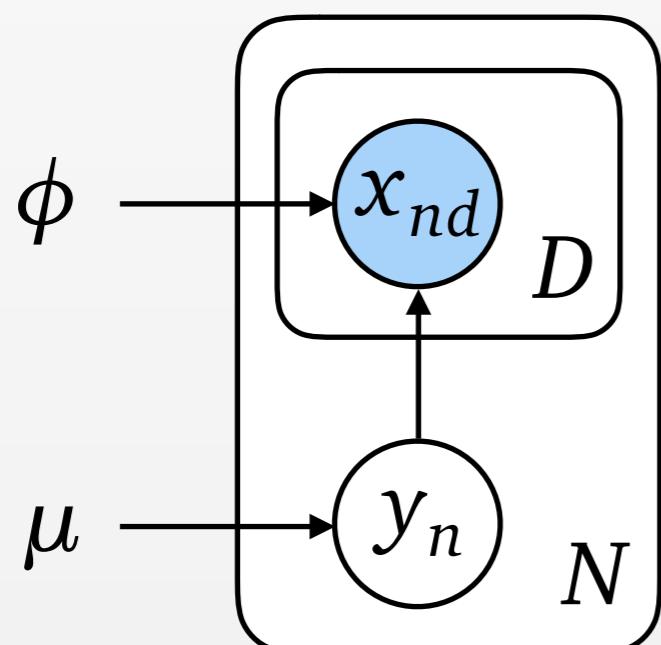
# Graphical Model: Naive Bayes

## Generative Model

$$y_n \sim \text{Bernoulli}(\mu) \quad n = 1, \dots, N$$

$$x_{nd} | y_n = k \sim \text{Bernoulli}(\phi_{kd}) \quad k = 0, 1 \quad d = 1, \dots, D$$

## Graphical Model



## Joint Distribution

$$p(x, y | \mu, \phi)$$

Goal: Predict Labels

$$p(y | x, \mu, \phi) = \frac{p(x | y, \phi)p(y | \mu)}{p(x | \mu, \phi)}$$

# Prediction with Bayesian Posterior

Posterior on Label (Spam=1, Not Spam=0)

$$p(y' | x', \mu, \phi) = \frac{p(x', y' | \mu, \phi)}{p(x' | \mu, \phi)}$$

Marginal Likelihood

$$p(x' | \mu, \phi) = \sum_{k=\{0,1\}} p(x', y' = k | \mu, \phi)$$

Joint

$$p(x', y' | \mu, \phi) = p(y' | \mu) \prod_{d=1} p(x'_d | y', \phi)$$

$$p(y' | \mu) = \mu^{y'} (1 - \mu)^{(1-y')}$$

$$p(x'_d | y' = k, \phi) = \phi_{kd}^{x'_d} (1 - \phi_{kd})^{1-x'_d}$$

# Parameter Estimation

Maximum Likelihood

$$\operatorname{argmax}_{\mu, \phi} \sum_{n=1}^N \log p(y_n, x_n | \mu, \phi)$$

Labeled Data

$$\{(x_1, y_1), \dots, (x_N, y_N)\}$$

Maximum A Posteriori

$$\operatorname{argmax}_{\mu, \phi} \log p(\mu, \phi) + \sum_{n=1}^N \log p(y_n, x_n | \mu, \phi)$$

Test-time Prediction

$$p(y' | x', \mu^*, \phi^*)$$

# Maximum Likelihood Estimation

Objective

$$\operatorname{argmax}_{\mu, \phi} \sum_{n=1}^N \log p(y_n, x_n \mid \mu, \phi)$$

Training Data (Labeled)

$$\{(x_1, y_1), \dots, (x_N, y_N)\}$$

Optimum Parameters

$$\mu = \frac{1}{N} \sum_{n=1}^N I[y_n = 1]$$

Interpretation: Fraction of Spam in training set

$$\phi_{kd} = \frac{1}{N_k} \sum_{n:y_n=k} I[x_{nd} = 1]$$

Interpretation: Fraction of *non-spam* ( $k=0$ ) and *spam* ( $k=1$ ) messages that contain term.

# Maximum Likelihood Estimation

Objective

$$\operatorname{argmax}_{\mu, \phi} \sum_{n=1}^N \log p(y_n, x_n \mid \mu, \phi)$$

Training Data (Labeled)

$$\{(x_1, y_1), \dots, (x_N, y_N)\}$$

Optimum Parameters

$$\mu = \frac{N_1^y}{N}$$

$$\phi_{kd} = \frac{N_{kd}^x}{N_k^y}$$

Problem:

What do you do for words not found in training set?

$$\sum_n I[x_{nd} = 1] = 0$$

$$\phi_{0d} = \phi_{1d} = 0$$

# Maximum Likelihood Estimation

Objective

$$\operatorname{argmax}_{\mu, \phi} \sum_{n=1}^N \log p(y_n, x_n \mid \mu, \phi)$$

Training Data (Labeled)

$$\{(x_1, y_1), \dots, (x_N, y_N)\}$$

Optimum Parameters

(with Laplace smoothing)

$$\mu = \frac{N_1^y + 1}{N + 2}$$

$$\phi_{kd} = \frac{N_{kd}^x + 1}{N_k^y + D}$$

Problem:

What do you do for words not found in training set?

$$\sum_n I[x_{nd} = 1] = 0$$

$$\phi_{0d} = \phi_{1d} = 0$$

# Naive Bayes with Priors

Generative Model

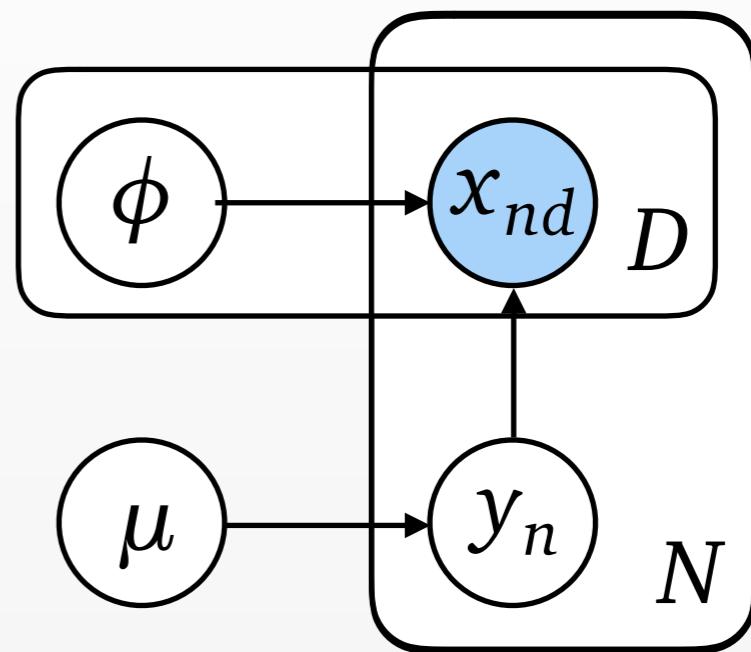
$$\mu \sim \text{Beta}(1, 1)$$

$$\phi_{kd} \sim \text{Beta}(1, 1)$$

$$y_n \sim \text{Bernoulli}(\mu)$$

$$x_{nd} \mid y_n = k \sim \text{Bernoulli}(\phi_{kd})$$

Graphical Model



Maximum A Posteriori

$$y' = \operatorname{argmax}_{y'} p(y' \mid x', \mu^*, \phi^*)$$

$$\mu^*, \phi^* = \operatorname{argmax}_{\mu, \phi} \log p(\mu, \phi) + \sum_{n=1}^N \log p(x_n, y_n \mid \mu, \phi)$$

# Exponential Families

# Exponential Family Distributions

$$p(x | \eta) = h(x) \exp \left[ \sum_i \eta_i t_i(x) - a(\eta) \right]$$

Diagram illustrating the components of the Exponential Family Distribution formula:

- Base Measure:**  $h(x)$  (represented by a purple dotted box)
- Log Normalizer:**  $a(\eta)$  (represented by a red dotted box)
- Natural Parameters:**  $\eta$  (represented by an orange arrow pointing up to the term  $\sum_i \eta_i t_i(x)$ )
- Sufficient Statistics:**  $t_i(x)$  (represented by an orange arrow pointing up to the term  $\sum_i \eta_i t_i(x)$ )

Product of terms:

$x$ -dependent,  $x\eta$ -dependent,  $\eta$ -dependent

# Example: Discrete Distribution

$$p(x \mid \eta) = h(x) \exp \left[ \sum_i \eta_i t_i(x) - a(\eta) \right]$$

$$p(x \mid \theta) = \prod_{k=1}^K \theta_k^{x_k} = \exp \left[ \sum_{k=1}^K \log(\theta_k) x_k \right]$$

$h(x) = 1$   
 $\eta_k = \log \theta_k$   
 $t_k(x) = x_k$   
 $a(\eta) = 0$

# Example: Gaussian Distribution

$$p(x \mid \eta) = h(x) \exp \left[ \sum_i \eta_i t_i(x) - a(\eta) \right]$$

$$h(x) = \frac{1}{\sqrt{2\pi}}$$

$$p(x \mid \mu, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left[ -\frac{1}{2} \frac{(x - \mu)^2}{\sigma^2} \right] \quad \eta = \left( \frac{\mu}{\sigma^2}, -\frac{1}{\sigma^2} \right)$$
$$t(x) = (x, x^2)$$

$$a(\eta) = \frac{\mu^2}{\sigma^2} + \log \sigma$$

# Exponential Families

[https://en.wikipedia.org/wiki/Exponential\\_family](https://en.wikipedia.org/wiki/Exponential_family)

Distribution	Parameter(s) $\theta$	Natural parameter(s) $\eta$	Inverse parameter mapping	Base measure $h(x)$	Sufficient statistic $T(x)$	Log-partition $A(\eta)$	Log-partition $A(\theta)$
Bernoulli distribution	$p$	$\ln \frac{p}{1-p}$ • This is the logit function.	$\frac{1}{1+e^{-\eta}} = \frac{e^\eta}{1+e^\eta}$ • This is the logistic function.	1	$x$	$\ln(1+e^\eta)$	$-\ln(1-p)$
binomial distribution with known number of trials $n$	$p$	$\ln \frac{p}{1-p}$	$\frac{1}{1+e^{-\eta}} = \frac{e^\eta}{1+e^\eta}$	$\binom{n}{x}$	$x$	$n \ln(1+e^\eta)$	$-n \ln(1-p)$
Poisson distribution	$\lambda$	$\ln \lambda$	$e^\eta$	$\frac{1}{x!}$	$x$	$e^\eta$	$\lambda$
negative binomial distribution with known number of failures $r$	$p$	$\ln p$	$e^\eta$	$\binom{x+r-1}{x} x$		$-r \ln(1-e^\eta)$	$-r \ln(1-p)$
exponential distribution	$\lambda$	$-\lambda$	$-\eta$	1	$x$	$-\ln(-\eta)$	$-\ln \lambda$
Pareto distribution with known minimum value $x_m$	$\alpha$	$-\alpha - 1$	$-1 - \eta$	1	$\ln x$	$-\ln(-1 - \eta) + (1 + \eta) \ln x_m$	$-\ln \alpha - \alpha \ln x_m$

Most of the “normally” used distributions:  
 Normal, Gamma, Dirichlet, Discrete, Poisson, Cauchy, etc.