Log-Linear Models a.k.a. Logistic Regression a.k.a. Maximum Entropy Models

Natural Language Processing CS 4120/6120—Spring 2015 Northeastern University

David Smith (some slides from Jason Eisner and Dan Klein)

Probability is Useful

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 - Bayesian smoothing: $\max p(\theta|\text{data}) = \max p(\theta, \text{data}) = p(\theta)p(\text{data}|\theta)$

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- p(...) has to capture our intuitions about the ling. data



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 - Possible parses (or whatever) have scores.
 - Pick the one with the best score.
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An Alternative Tradition

- Old A
 - Pos
 - Pic
 - Ho

- Lear" pena
- Total
 - Ca

Exposé at 9

Probabilistic Revolution Not Really a Revolution, Critics Say

Log-probabilities no more than scores in disguise

"We're just adding stuff up like the old corrupt regime did," admits spokesperson



uses and nce. [©]

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- *Note:* Today we'll use +logprob not –logprob: i.e., bigger weights are **better**.

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- PCFG: log p(NP VP | S) + log p(Papa | NP) + log p(VP PP | VP) ...
 - Can regard any linguistic object as a collection of features (here, tree = a collection of context-free rules)
 - Weight of the object = total weight of features
 - Our weights have always been conditional log-probs (≤ 0)
 - but that is going to change in a few minutes!
- HMM tagging: ... + log p(t7 | t5, t6) + log p(w7 | t7) + ...
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but here are the emails with both features – only 25x!

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Sparn time

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Contains a dollar amount under \$100 -.15 | -3.3

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- Well, maybe we can add up scores and <u>pretend</u> like we got a log probability: log p(feats | spam) = 5.77

- Naïve Bayes needs overlapping but independent features
- But not clear how to restructure these features like that:

```
total: 5.77
       Contains Buy
+0.2
        Contains supercalifragilistic
+1
       Contains a dollar amount under $100
        Contains an imperative sentence
        Reading level = 7^{th} grade

    Mentions money (use word classes and/or regexp to detect this)
```

- Boy, we'd like to be able to throw all that useful stuff in without worrying about feature overlap/independence.
- Well, maybe we can add up scores and <u>pretend</u> like we got a log probability: log p(feats | spam) = 5.77Oops, then p(feats | spam) = exp 5.77 = 320.5

Renormalize by 1/Z to get a Log-Linear Model

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Renormalize by 1/Z to get a p(feats | spam) = $\exp 5.77 = 320.5$ and $\sup_{\text{everything to 1!}} \sup_{\text{and sums to 1!}} \sup_{\text{span}} \sup_{$

Renormalize by 1/Z to get a scale down so. **Log-Linear Model**

- everything < 1 and sums to 1! • p(feats | spam) = exp 5.77 = 320.5)
- $p(m \mid spam) = (1/Z(\lambda)) \exp \sum_{i} \lambda_{i} f_{i}(m)$ where m is the email message λ_i is weight of feature i $f_i(m) \in \{0,1\}$ according to whether m has feature i
 - More generally, allow $f_i(m) = count or strength of feature.$
 - $1/Z(\lambda)$ is a normalizing factor making $\sum_{m} p(m \mid spam)=1$ (summed over all possible messages m! hard to find!)

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 - (summed over all possible messages m! hard to find!)
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- Why is it called "log-linear"?

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- But: p(m_i | c_i) for a given λ requires Z(λ): hard!

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- But we can fix this ...

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 - spam and Contains Buy ←old spam model's weight for "contains Buy"
 - spam and Contains supercalifragilistic
 - ____
 - ling ← weight of this feature is log p(ling) + a constant
 - ling and Contains Buy ←old ling model's weight for "contains Buy"
 - ling and Contains supercalifragilistic
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17

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 - Easy to compute now ...
 - $\prod_i p(c_i \mid m_i)$ is still convex, so easy to maximize too

Generative vs. Conditional

- What is the most likely label for a given input?
- How likely is a given label for a given input?
- What is the most likely input value?
- How likely is a given input value?
- How likely is a given input value with a given label?
- What is the most likely label for an input that might have one of two values (but we don't know which)?

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- OUCH!

	Α	В	С	D	Е	F	G	Н	Ι	J
Buy	0.051	0.0025	0.029	0.0025	0.0025	0.0025	0.0025	0.0025	0.0025	0.0025
Other	0.499	0.0446	0.0446	0.0446	0.0446	0.0446	0.0446	0.0446	0.0446	0.0446

Column A sums to 0.55 ("55% of all messages are in class A")

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- Column A sums to 0.55
- Row Buy sums to 0.1 ("10% of all messages contain Buy")

	Α	В	С	D	Е	F	G	Н	Ι	J
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- Column A sums to 0.55
- Row Buy sums to 0.1
- (Buy, A) and (Buy, C) cells sum to 0.08 ("80% of the 10%")

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Entropy = $-.051 \log .051 - .0025 \log .0025 - .029 \log .029 - ...$

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Entropy = -.051 log .051 - .0025 log .0025 - .029 log .029 - ... Largest if probabilities are evenly distributed

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- Given these constraints, fill in cells "as equally as possible": maximize the entropy
- Now p(Buy, C) = .029 and $p(C \mid Buy) = .29$
- We got a compromise: p(C | Buy) < p(A | Buy) < .55</p>

Generalizing to More Features

	\$100/								
Othe	<u>r/</u>								
	A	В	С	D	E	F	G	H	
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 - So it is log-linear. In fact it is the same log-linear distribution that maximizes $\prod_j p(m_j, c_j)$ as before!
 - Gives another motivation for the log-linear approach.

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Log-linear form derivation

 Say we are given some constraints in the form of feature expectations:

$$\sum_{x} p(x) f_i(x) = \alpha_i$$

- In general, there may be many distributions p(x) that satisfy the constraints. Which one to pick?
- The one with maximum entropy (making fewest possible additional assumptions---Occum's Razor)
- This yields an optimization problem

$$\max H(p(x)) = -\sum_x p(x) \log p(x)$$
 Subject to
$$\sum_x p(x) f_i(x) = \alpha_i, \forall i \ \text{ and } \sum_x p(x) = 1$$

Log-linear form derivation

To solve the maxent problem, we use Lagrange multipliers:

$$L = -\sum_{\mathbf{x}} p(\mathbf{x}) \log p(\mathbf{x}) - \sum_{i} \theta_{i} \left(\sum_{\mathbf{x}} p(\mathbf{x}) f_{i}(\mathbf{x}) - \alpha_{i} \right) - \mu \left(\sum_{\mathbf{x}} p(\mathbf{x}) - 1 \right)$$

$$\frac{\partial L}{\partial p(\mathbf{x})} = 1 + \log p(\mathbf{x}) - \sum_{i} \theta_{i} f_{i}(\mathbf{x}) - \mu$$

$$p^{*}(\mathbf{x}) = e^{\mu - 1} \exp \left\{ \sum_{i} \theta_{i} f_{i}(\mathbf{x}) \right\}$$

$$Z(\theta) = e^{1 - \mu} = \sum_{\mathbf{x}} \exp \left\{ \sum_{i} \theta_{i} f_{i}(\mathbf{x}) \right\}$$

$$p(\mathbf{x}|\theta) = \frac{1}{Z(\theta)} \exp \left\{ \sum_{i} \theta_{i} f_{i}(\mathbf{x}) \right\}$$

- So feature constraints + maxent implies exponential family.
- Problem is convex, so solution is unique.

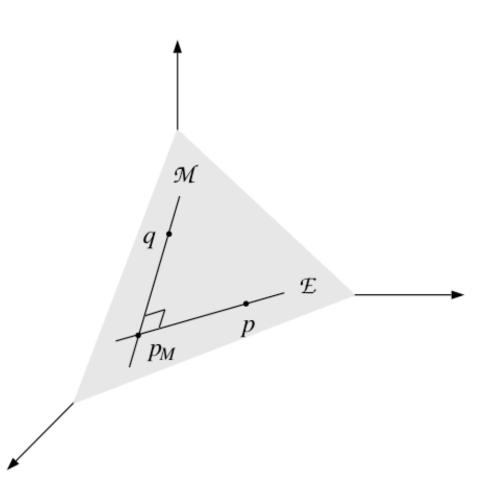
MaxEnt = Max Likelihood

Define two submanifolds on the probability simplex $p(\mathbf{x})$.

The first is \mathcal{E} , the set of all exponential family distributions based on a particular set of features $f_i(\mathbf{x})$.

The second is \mathcal{M} , the set of all distributions that satisfy the feature expectation constraints.

They intersect at a single distribution p_M , the maxent, maximum likelihood





Exponential Model Likelihood

- Maximum Likelihood (Conditional) Models :
 - Given a model form, choose values of parameters to maximize the (conditional) likelihood of the data.
- Exponential model form, for a data set (C,D):

$$\log P(C \mid D, \lambda) = \sum_{(c,d) \in (C,D)} \log P(c \mid d, \lambda) = \sum_{(c,d) \in (C,D)} \log \frac{\exp \sum_{i} \lambda_{i} f_{i}(c,d)}{\sum_{c'} \exp \sum_{i} \lambda_{i} f_{i}(c',d)}$$



Building a Maxent Model

- Define features (indicator functions) over data points.
 - Features represent sets of data points which are distinctive enough to deserve model parameters.
 - Usually features are added incrementally to "target" errors.
- For any given feature weights, we want to be able to calculate:
 - Data (conditional) likelihood
 - Derivative of the likelihood wrt each feature weight
 - Use expectations of each feature according to the model
- Find the optimum feature weights (next part).



The Likelihood Value

 The (log) conditional likelihood is a function of the iid data (C,D) and the parameters λ:

$$\log P(C \mid D, \lambda) = \log \prod_{(c,d) \in (C,D)} P(c \mid d, \lambda) = \sum_{(c,d) \in (C,D)} \log P(c \mid d, \lambda)$$

If there aren't many values of c, it's easy to calculate:

$$\log P(C \mid D, \lambda) = \sum_{(c,d) \in (C,D)} \log \frac{\exp \sum_{i} \lambda_{i} f_{i}(c,d)}{\sum_{c'} \exp \sum_{i} \lambda_{i} f_{i}(c,d)}$$

We can separate this into two components:

$$\log P(C \mid D, \lambda) = \sum_{(c,d) \in (C,D)} \log \exp \sum_{i} \lambda_{i} f_{i}(c,d) - \sum_{(c,d) \in (C,D)} \log \sum_{c'} \exp \sum_{i} \lambda_{i} f_{i}(c',d)$$

$$\log P(C \mid D, \lambda) = N(\lambda) - M(\lambda)$$

The derivative is the difference between the derivatives of each component



The Derivative I: Numerator

$$\frac{\partial N(\lambda)}{\partial \lambda_{i}} = \frac{\partial \sum_{(c,d) \in (C,D)} \log \exp \sum_{i} \lambda_{ci} f_{i}(c,d)}{\partial \lambda_{i}} = \frac{\partial \sum_{(c,d) \in (C,D)} \sum_{i} \lambda_{i} f_{i}(c,d)}{\partial \lambda_{i}}$$

$$= \sum_{(c,d) \in (C,D)} \frac{\partial \sum_{i} \lambda_{i} f_{i}(c,d)}{\partial \lambda_{i}}$$

$$= \sum_{(c,d) \in (C,D)} f_{i}(c,d)$$

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Derivative of the numerator is: the empirical count(f_i , c)



The Derivative II: Denominator

$$\frac{\partial M(\lambda)}{\partial \lambda_{i}} = \frac{\partial \sum_{(c,d) \in (C,D)} \log \sum_{c'} \exp \sum_{i} \lambda_{i} f_{i}(c',d)}{\partial \lambda_{i}}$$

$$= \sum_{(c,d) \in (C,D)} \frac{1}{\sum_{c''} \exp \sum_{i} \lambda_{i} f_{i}(c'',d)} \frac{\partial \sum_{c'} \exp \sum_{i} \lambda_{i} f_{i}(c',d)}{\partial \lambda_{i}}$$

$$= \sum_{(c,d) \in (C,D)} \frac{1}{\sum_{c''} \exp \sum_{i} \lambda_{i} f_{i}(c'',d)} \sum_{c'} \frac{\exp \sum_{i} \lambda_{i} f_{i}(c',d)}{1} \frac{\partial \sum_{i} \lambda_{i} f_{i}(c',d)}{\partial \lambda_{i}}$$

$$= \sum_{(c,d) \in (C,D)} \sum_{c'} \frac{\exp \sum_{i} \lambda_{i} f_{i}(c',d)}{\sum_{c''} \exp \sum_{i} \lambda_{i} f_{i}(c',d)} \frac{\partial \sum_{i} \lambda_{i} f_{i}(c',d)}{\partial \lambda_{i}}$$

$$= \sum_{(c,d) \in (C,D)} \sum_{c'} \frac{P(c'|d,\lambda) f_{i}(c',d)}{\sum_{c''} \exp \sum_{i} \lambda_{i} f_{i}(c',d)} = \text{predicted count}(f_{i},\lambda)$$



The Derivative III

$$\frac{\partial \log P(C \mid D, \lambda)}{\partial \lambda_i} = \operatorname{actual count}(f_i, C) - \operatorname{predicted count}(f_i, \lambda)$$

- The optimum parameters are the ones for which each feature's predicted expectation equals its empirical expectation. The optimum distribution is:
 - Always unique (but parameters may not be unique)
 - Always exists (if features counts are from actual data).
- Features can have high model expectations (predicted counts) either because they have large weights or because they occur with other features which have large weights.



We have a function to optimize:

$$\log P(C \mid D, \lambda) = \sum_{(c,d) \in (C,D)} \log \frac{\exp \sum_{i} \lambda_{i} f_{i}(c,d)}{\sum_{c'} \exp \sum_{i} \lambda_{i} f_{i}(c,d)}$$

We know the function's derivatives:

$$\partial \log P(C \mid D, \lambda) / \partial \lambda_i = \operatorname{actual count}(f_i, C) - \operatorname{predicted count}(f_i, \lambda)$$

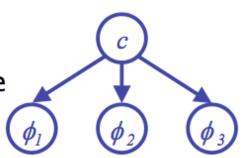
Perfect situation for general optimization (Part II)

By gradient ascent or conjugate gradient.



Comparison to Naïve-Bayes

- Naïve-Bayes is another tool for classification:
 - We have a bunch of random variables (data features) which we would like to use to predict another variable (the class):



The Naïve-Bayes likelihood over classes is:

$$P(c \mid d, \lambda) = \frac{P(c) \prod_{i} P(\phi_{i} \mid c)}{\sum_{c'} P(c') \prod_{i} P(\phi_{i} \mid c')} \longrightarrow \frac{\exp\left[\log P(c) + \sum_{i} \log P(\phi_{i} \mid c)\right]}{\sum_{c'} \exp\left[\log P(c') + \sum_{i} \log P(\phi_{i} \mid c')\right]}$$

Naïve-Bayes is just an exponential model.

$$\sum_{i} \lambda_{ic} f_{ic}(d,c)$$
 $\sum_{i} \exp \left[\sum_{i} \lambda_{ic'} f_{ic'}(d,c')\right]$



Comparison to Naïve-Bayes

The primary differences between Naïve-Bayes and maxent models are:

Naïve-Bayes

Trained to maximize joint likelihood of data and classes.

Features assumed to supply independent evidence.

Feature weights can be set independently.

Features must be of the conjunctive $\Phi(d) \wedge c = c_i$ form.

Maxent

Trained to maximize the conditional likelihood of classes.

Features weights take feature dependence into account.

Feature weights must be mutually estimated.

Features need not be of the conjunctive form (but usually are).

Overfitting

If we have too many features, we can choose weights to model the training data perfectly.

If we have a feature that only appears in spam training, not ling training, it will get weight ∞ to maximize p(spam | feature) at 1.

- These behaviors overfit the training data.
- Will probably do poorly on test data.

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- 3. Smooth the observed feature counts.
- 4. Smooth the weights by using a prior.
 - max $p(\lambda|data) = max p(\lambda, data) = p(\lambda)p(data|\lambda)$
 - decree $p(\lambda)$ to be high when most weights close to 0



Smoothing: Priors (MAP)

- What if we had a prior expectation that parameter values wouldn't be very large?
- We could then balance evidence suggesting large parameters (or infinite) against our prior.
- The evidence would never totally defeat the prior, and parameters would be smoothed (and kept finite!).
- We can do this explicitly by changing the optimization objective to maximum posterior likelihood:

$$\log P(C, \lambda \mid D) = \log P(\lambda) + \log P(C \mid D, \lambda)$$

Posterior Prior Evidence

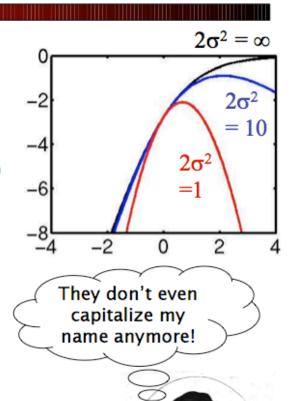


Smoothing: Priors

- Gaussian, or quadratic, priors:
 - Intuition: parameters shouldn't be large.
 - Formalization: prior expectation that each parameter will be distributed according to a gaussian with mean μ and variance σ^2 .

$$P(\lambda_i) = \frac{1}{\sigma_i \sqrt{2\pi}} \exp\left(-\frac{(\lambda_i - \mu_i)^2}{2\sigma_i^2}\right)$$

- Penalizes parameters for drifting to far from their mean prior value (usually µ=0).
- $2\sigma^2=1$ works surprisingly well.



Recipe for a Conditional MaxEnt Classifier

1. Gather constraints from training data:

$$\alpha_{iy} = \tilde{E}[f_{iy}] = \sum_{x_j, y_j \in D} f_{iy}(x_j, y_j)$$

- 2. Initialize all parameters to zero.
- 3. Classify training data with current parameters. Calculate expectations. $E_{\Theta}[f_{iy}] = \sum p_{\Theta}(y'|x_j)f_{iy}(x_j,y')$

 $x_i \in D y'$

- 4. Gradient is $\tilde{E}[f_{iy}] E_{\Theta}[f_{iy}]$
- 5. Take a step in the direction of the gradient
- 6. Until convergence, return to step 3.